## Online Appendix - Not For Publication

## Strategic Patient Discharge: The Case of Long-Term Care Hospitals

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## A Complete Summary Statistics

Table A1: Summary Statistics for All Patients (2004-2013)

| Variable | Mean | Std. Dev. |
| :--- | :---: | :---: |
| Length of Stay | 28.766 | 41.844 |
| Released on or after magic day | 0.681 | 0.466 |
| Total Payment (\$) | $31,933.43$ | $24,332.54$ |
| Amount Paid by Medicare (\$) | $31,814.61$ | $26,883.69$ |
| Estimated Costs (\$) | $37,578.69$ | $37,022.04$ |
| Portion Discharged Alive | 0.861 | 0.346 |
| Portion Discharged Dead | 0.139 | 0.346 |
| Portion Discharged to Home Care | 0.34 | 0.474 |
| Portion Discharged to Hospital | 0.123 | 0.329 |
| Portion Discharged to Nursing Facility | 0.391 | 0.488 |
| Admission Type: Emergency | 0.011 | 0.104 |
| Admission Type: Urgent | 0.198 | 0.398 |
| Admission Type: Elective | 0.785 | 0.411 |
| Admission Type: Other | 0.006 | 0.079 |
| Admission Source: Community | 0.186 | 0.389 |
| Admission Source: Nursing Facility | 0.025 | 0.155 |
| Admission Source: General Hospital | 0.777 | 0.416 |
| Admission Source: Other Source | 0.007 | 0.085 |
| Male | 0.484 | 0.5 |
| White | 0.729 | 0.445 |
| African-American | 0.202 | 0.401 |
| Asian | 0.012 | 0.111 |
| Hispanic | 0.033 | 0.18 |
| Age less than 25 | 0.001 | 0.038 |
| Ave between 25 and 44 | 0.039 | 0.193 |
| Age between 45 and 64 | 0.191 | 0.393 |
| Age between 65 and 74 | 0.305 | 0.46 |
| Age between 75 and 84 | 0.301 | 0.459 |
| Age over 85 | 0.164 | 0.37 |
| $N=1,452,287$ |  |  |

Table A2: Summary Statistics for DRG 207 Patients (2004-2013)

| Variable | Mean | Std. Dev. |
| :--- | :---: | :---: |
| Length of Stay | 38.06 | 40.24 |
| Released on or after magic day | 0.672 | 0.47 |
| Total Payment (\$) | $57,609.66$ | $33,061.67$ |
| Amount Paid by Medicare (\$) | $57,536.17$ | $37,143.23$ |
| Estimated Costs (\$) | $67,061.07$ | $51,780.64$ |
| Portion Discharged Alive | 0.736 | 0.441 |
| Portion Discharged Dead | 0.264 | 0.441 |
| Portion Discharged to Home Care | 0.132 | 0.338 |
| Portion Discharged to Hospital | 0.166 | 0.372 |
| Portion Discharged to Nursing Facility | 0.437 | 0.496 |
| Admission Type: Emergency | 0.011 | 0.105 |
| Admission Type: Urgent | 0.202 | 0.402 |
| Admission Type: Elective | 0.781 | 0.414 |
| Admission Type: Other | 0.006 | 0.076 |
| Admission Source: Community | 0.122 | 0.327 |
| Admission Source: Nursing Facility | 0.013 | 0.115 |
| Admission Source: General Hospital | 0.857 | 0.35 |
| Admission Source: Other Source | 0.003 | 0.054 |
| Male | 0.502 | 0.5 |
| White | 0.745 | 0.436 |
| African-American | 0.192 | 0.394 |
| Asian | 0.015 | 0.122 |
| Hispanic | 0.024 | 0.154 |
| Age less than 25 | 0.002 | 0.04 |
| Ave between 25 and 44 | 0.03 | 0.17 |
| Age between 45 and 64 | 0.187 | 0.39 |
| Age between 65 and 74 | 0.355 | 0.478 |
| Age between 75 and 84 | 0.32 | 0.466 |
| Age over 85 | 0.107 | 0.309 |
| $N=170,365$ |  |  |

Table A3: Share of discharges on the magic day and the preceding day

| Comparison Set | Day before <br> magic day | Magic Day | Ratio | P-value $^{1}$ | Diff-in- <br> Ratios | P-value $^{2}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Home | 0.017 | 0.103 | 6.06 | 0.000 |  |  |
| Nursing Facility | 0.009 | 0.076 | 8.44 | 0.000 | -2.38 | 0.010 |
| Acute Care Hospital | 0.016 | 0.024 | 1.5 | 0.001 | 4.56 | 0.000 |
| Death | 0.018 | 0.019 | 1.06 | 0.517 | 5.01 | 0.000 |
| 2004 | 0.016 | 0.036 | 2.25 | 0.000 | 3.19 | 0.000 |
| 2013 | 0.016 | 0.087 | 5.44 | 0.000 |  |  |
| For-profit | 0.010 | 0.092 | 9.20 | 0.000 | 4.60 | 0.000 |
| Non-profit | 0.015 | 0.069 | 4.60 | 0.000 |  |  |
| Select or Kindred | 0.010 | 0.089 | 8.91 | 0.000 | 3.29 | 0.000 |
| Other | 0.013 | 0.073 | 5.62 | 0.000 |  |  |
| Before Acquisition | 0.014 | 0.087 | 6.21 | 0.000 | 8.93 | 0.000 |
| After Acquisition | 0.007 | 0.106 | 15.14 | 0.000 |  |  |
| Co-located | 0.012 | 0.101 | 8.42 | 0.000 | 1.78 | 0.074 |
| Not Co-located | 0.011 | 0.073 | 6.64 | 0.000 |  |  |

Note: P-values from Wald tests of nonlinear hypotheses. Difference-in-ratios for nursing facility, acute-care hospital, and death discharges are all with respect to home discharges. Except for the discharge destination rows, the statistics include hospital stays ending in discharge to home or nursing facility care.
${ }^{1} \mathrm{P}$-value under the null hypothesis that the ratio is equal to one.
${ }^{2} \mathrm{P}$-value under the null hypothesis that the difference-in-ratios equals zero.
Table A4: Summary Statistics for Nine Most Common DRGs

|  | DRG |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 177 | 189 | 190 | 193 | 207 | 539 | 592 | 871 | 949 | Pooled |
| Mean length of stay | 25.3 | 26.4 | 21.0 | 22.3 | 42.4 | 33.2 | 30.4 | 26.0 | 24.2 | 30.0 |
| Standard deviation | (12.6) | (20.2) | (10.0) | (12.2) | (24.1) | (15.0) | (16.6) | (14.1) | (18.1) | (19.6) |
|  |  | Paym | and | st Estim | mates (in | \$) |  |  |  |  |
| Mean daily payments | 1,186 | 1,245 | 1,139 | 1,124 | 1,639 | 1,013 | 974 | 1,100 | 981 | 1,249 |
| Mean full payments | 33,466 | 39,929 | 27,289 | 28,401 | 7,8749 | 36,334 | 33,594 | 33,307 | 27,153 | 44,626 |
| Mean magic day payments | 9,116 | 1,3264 | 8,087 | 7,846 | 33,562 | 8,857 | 11,765 | 12,356 | 9,488 | 16,308 |
| Mean daily cost est. | 1,267 | 1,341 | 1,191 | 1,184 | 1,689 | 1,081 | 1,026 | 1,179 | 1,098 | 1,319 |
|  |  |  | Disch | ge Ty |  |  |  |  |  |  |
| Discharged alive | 0.84 | 0.83 | 0.89 | 0.85 | 0.73 | 0.94 | 0.87 | 0.83 | 0.95 | 0.82 |
| Discharged to home | 0.29 | 0.29 | 0.54 | 0.39 | 0.13 | 0.41 | 0.30 | 0.28 | 0.41 | 0.28 |
| Discharged to hospital | 0.09 | 0.11 | 0.08 | 0.09 | 0.16 | 0.13 | 0.12 | 0.10 | 0.14 | 0.12 |
| Discharged to nursing facility | 0.46 | 0.42 | 0.26 | 0.37 | 0.44 | 0.39 | 0.44 | 0.44 | 0.39 | 0.40 |
|  |  |  |  | H Type |  |  |  |  |  |  |
| For-profit, HwH | 0.14 | 0.17 | 0.14 | 0.12 | 0.23 | 0.17 | 0.18 | 0.13 | 0.17 | 0.17 |
| For-profit, standalone | 0.61 | 0.56 | 0.57 | 0.63 | 0.50 | 0.54 | 0.58 | 0.67 | 0.56 | 0.57 |
| Non-profit, HwH | 0.07 | 0.08 | 0.05 | 0.06 | 0.09 | 0.10 | 0.07 | 0.06 | 0.03 | 0.07 |
| Non-profit, standalone | 0.17 | 0.18 | 0.23 | 0.18 | 0.18 | 0.18 | 0.18 | 0.16 | 0.14 | 0.24 |
| $N^{1}$ | 38,318 | 71,563 | 28,139 | 26,492 | 90,755 | 18,923 | 36,669 | 50,494 | 16,160 | 377,513 |

Table A5: Per-diem estimates (in \$)

|  | Mean | 25th <br> Percentile | Median | 75th <br> Percentile |
| :---: | :---: | :---: | :---: | :---: |
| Panel A: Per-diem rate |  |  |  |  |
| Overall | 1,249 | 1,050 | 1,195 | 1,414 |
| For-profit, HwH | 1,235 | 1,005 | 1,179 | 1,480 |
| For-profit, standalone | 1,228 | 1,043 | 1,178 | 1,368 |
| Non-profit, HwH | 1,280 | 1,055 | 1,220 | 1,503 |
| Non-profit, standalone | 1,317 | 1,117 | 1,257 | 1,507 |
| Select | 1,250 | 1,024 | 1,206 | 1,482 |
| Kindred | 1,232 | 1,049 | 1,187 | 1,377 |
| Other | 1,257 | 1,058 | 1,198 | 1,405 |
| Panel B: Full LTCH PPS payment |  |  |  |  |
| Overall | 44,626 | 30,938 | 35,155 | 61,702 |
| For-profit, HwH | 46,876 | 30,517 | 35,195 | 72,845 |
| For-profit, standalone | 43,817 | 31,318 | 35,208 | 43,111 |
| Non-profit, HwH | 44,177 | 33,746 | 33,746 | 68,396 |
| Non-profit, standalone | 45,256 | 30,853 | 35,302 | 63,558 |
| Select | 47,480 | 31,310 | 35,577 | 73,571 |
| Kindred | 46,358 | 33,097 | 36,889 | 59,953 |
| Other | 42,661 | 30,092 | 34,059 | 42,658 |
| Panel C: Magic day payments |  |  |  |  |
| Overall | 16,308 | 8,742 | 12,450 | 22,710 |
| For-profit, HwH | 17,763 | 8,965 | 13,529 | 29,478 |
| For-profit, standalone | 16,351 | 9,209 | 12,630 | 20,749 |
| Non-profit, HwH | 14,437 | 7,018 | 11,352 | 23,111 |
| Non-profit, standalone | 15,536 | 7,918 | 11,127 | 24,193 |
| Select | 18,114 | 9,592 | 13,742 | 30,162 |
| Kindred | 18,448 | 10,666 | 14,591 | 25,715 |
| Other | 14,555 | 7,763 | 11,234 | 18,895 |
| $N=377,513$ |  |  |  |  |

Table A6: Average daily cost estimates (in \$)

|  | Mean | 25th <br> Percentile | Median | 75th <br> Percentile |
| :--- | :---: | :---: | :---: | :---: |
| Overall | 1,319 | 1,075 | 1,280 | 1,526 |
| For-profit, HwH | 1,266 | 1,003 | 1,237 | 1,501 |
| For-profit, standalone | 1,300 | 1,078 | 1,267 | 1,488 |
| Non-profit, HwH | 1,398 | 1,100 | 1,365 | 1,640 |
| Non-profit, standalone | 1,401 | 1,135 | 1,372 | 1,631 |
| Select | 1,279 | 1,028 | 1,267 | 1,497 |
| Kindred | 1,293 | 1,078 | 1,253 | 1,487 |
| Other | 1,348 | 1,089 | 1,301 | 1,558 |
| $N=377,513$ |  |  |  |  |

## B Other DRGs

While our paper at times focuses on DRG 207, in this appendix we extend the analysis to other
DRGs, summarized above in Appendix A. Our structural estimation uses the nine most common DRGs in order to increase the variation in magic days in the data. Table A7 describes each of these DRGs. Figure A1 plots discharge patterns for the next three most common DRGs after DRG 207 in 2004 and 2013, along with their respective SSO thresholds. Figure A2 plots realized Medicare payments and discharge patterns that suggest other DRGs have similar discharge practices.

Table A7: DRG Descriptions

| DRG | Description |
| :--- | :--- |
| 177 | Respiratory infections and inflammations with <br> major complicating conditions |
| 189 | Pulmonary edema and respiratory failure |
| 190 | Chronic obstructive pulmonary disease with major <br> complicating conditions |
| 193 | Simple pneumonia and pleurisy with major complicating <br> conditions |
| 207 | Respiratory system diagnosis with ventilator support of <br> over 96 hours |
| 539 | Osteomyelitis with major complicating conditions |
| 592 | Skin ulcers |
| 871 | Septicemia without mechanical ventilation of over <br> $96 ~ h o u r s ~ w i t h ~ m a j o r ~ c o m p l i c a t i n g ~ c o n d i t i o n s ~$ |
| 949 | Aftercare with complication conditions or major <br> complicating conditions |


(a) Discharge practices for DRG 189 in 2004

(c) Discharge practices for DRG 871 in 2004

(e) Discharge practices for DRG 177 in 2004

(b) Discharge practices for DRG 189 in 2013

(d) Discharge practices for DRG 871 in 2013

(f) Discharge practices for DRG 177 in 2013

Figure A1: Discharge timing across DRGs and years


Figure A2: Costs, payoffs and discharge patterns for other DRGs

## C Payment Policy Details \& Example

Medicare calculates the PPS by starting with an LTCH Standard Federal Rate, or LTCH base rate, which was $\$ 39,794.95$ in FY2010. Two adjustments are then applied to this base rate. The first is a hospital wage index adjustment that incorporates geographic differences in costs due to health-sector wages. The second is a Medicare severity long-term care diagnosis related
group (MT-LTC-DRG) adjustment. The MT-LTC-DRG weight adjusts the payment to account for patient diagnoses (principal and secondary), procedures, age, sex, and discharge status based on the expected relative costliness of patients in each group. The final adjusted amount is known as the full LTCH payment.
For short stays, Medicare pays LTCHs the least of the following:

1. The full MS-LTC-DRG payment, or
2. 100 percent of the cost of the case, or
3. 120 percent of the MS-LTC-DRG specific per-diem amount multiplied by the length of stay, or
4. A blend of the inpatient MS-DRG amount and 120 percent of the LTCH per-diem amount, where the portion coming from the LTCH per-diem amount increases with the length of stay.

Starting in calendar year 2013 there is also a "very short stay outlier" payment. Cases with stays less than or equal to the IPPS average length of stay are reimbursed at weakly lower rates than SSOs. These payments are set to the least of the four possibilities enumerated in the SSO case above but replace the blended case with just the inpatient MS-DRG amount ${ }^{43}$

## Full MS-LTC-DRG payment

## Example of Full LTCH-PPS Payment in 2010, DRG 207

LTCH Base Rate
\$39,794.95
Labor-related portion of base rate

$$
\$ 39,794.95 \times 0.75779=\$ 30,156.22
$$

Non-labor related portion of base rate

$$
\$ 39,794.95 \times 0.24221=\$ 9,638.73
$$

Labor-related portion adjusted for wage index (CBSA 16974) $\quad \$ 30,156.22 \times 1.0471=\$ 31,576.57$
Wage-adjusted LTCH Base Rate
\$41,215.31
MS-LTC-DRG 207 Relative Weight
2.0288

Total Adjusted Federal Prospective Payment
$\$ 41,215.31 \times 2.0288=\$ 83,617.62$
For more examples of computing full LTCH-PPS payments, see CMS document:
https://www.cms.gov/Medicare/Medicare-Fee-for-Service-Payment/ LongTermCareHospitalPPS/Downloads/LTCH_sso_ex_2007and2008.zip.

[^0]
## 100 percent of cost of case

$$
\text { cost of case }=(\text { covered charges }) \times(\text { cost-to-charge ratio })
$$

The hospital-specific cost-to-charge ratio is just what it sounds like. It is calculated for each hospital using cost data from the most recent cost report submitted from that hospital. Hospital CCR has two parts: operative CCR (total Medicare operating costs / total Medicare operating charges) and capital CCR (total Medicare capital costs / total Medicare capital charges). The CCR for each year is published in the LTCH Impact Files in August before the year begins and is based on most recent historical Medicare cost reports which are required on an annual basis.

$$
\begin{gathered}
\$ 45,501.00 \times 0.311=\$ 14,150.81=\text { Estimated Cost } \\
* \text { Assumes covered charges }=\$ 45,501.00 \text { and hospital CCR }=0.311
\end{gathered}
$$

## 120 percent of per-diem amount

*LTC-DRG average length of stay: 26.6 days. This case assumes an 8 day length of stay.

$$
\begin{aligned}
\text { MS-LTC-DRG per diem } & =\text { Full LTC-DRG Payment } / \text { Average Length of Stay of the LTC-DRG } \\
& =\$ 45,060.70 / 26.6 \text { days } \\
& =\$ 1,698.34 \text { per day } \\
120 \text { percent of per-diem amount } & =\$ 1,698.34 \times 8 \text { days } \times 1.2 \\
& =\$ 16,304.06
\end{aligned}
$$

## Blend Alternative

Computing the IPPS payment is considerably more involved, so for this example we simply assume it is $\$ 24,442.17$. The portion coming from the 120 percent of LTCH per diem is: $\frac{\text { length of stay }}{\text { SSO threshold }}=\frac{8}{22.2}=0.36$. The rest comes from the inpatient comparable per-diem amount that, after a complex series of calculations, is $\$ 24,442.17$. The blended amount is then:

$$
0.36 \times \$ 16,304.06+0.64 \times \$ 24442.17=\$ 21,512.45
$$

Since the "100 percent of cost" amount is the least, the law indicates that it is the will be paid out.
After our data period, CMS has continued to make changes to the LTCH-PPS. In FY 2018, the "very short stay outlier" payment was eliminated. In addition, CMS began calculating payments for discharges prior to the SSO threshold according to the "blended" formula described in option 4 above rather than the lowest payment of all four options. This change, which effectively raised payments for discharges prior to the SSO threshold, was implemented in part due to concerns that hospitals were delaying discharges in response to the SSO policy. ${ }^{44}$ However, the revised policy still results in a discontinuity in payments at the SSO threshold.

[^1]
## D Probit Model: Coefficient Estimates and other DRGs

Table A8 contains the estimated coefficients from the probit model for DRG 207. Table A9 presents the estimated marginal effects of the baseline probit model for other DRGs. Table A10 shows (a sample of) the estimated probit coefficients for the interacted models for DRG 207.

Table A8: Probit Estimates for DRG 207

|  | Coefficients | Std. Err. |
| :--- | :---: | :---: |
| Days relative to magic day $(\lambda s)$ |  |  |
| -14 | 0 | (Omitted group) |
| -13 | -0.021 | $(0.022)$ |
| -12 | 0.068 | $(0.026)$ |
| -11 | 0.103 | $(0.029)$ |
| -10 | 0.193 | $(0.032)$ |
| -9 | 0.333 | $(0.036)$ |
| -8 | 0.446 | $(0.041)$ |
| -7 | 0.497 | $(0.046)$ |
| -6 | 0.482 | $(0.051)$ |
| -5 | 0.486 | $(0.053)$ |
| -4 | 0.514 | $(0.062)$ |
| -3 | 0.522 | $(0.066)$ |
| -2 | 0.568 | $(0.070)$ |
| -1 | 0.665 | $(0.075)$ |
| 0 | 1.601 | $(0.080)$ |
| 1 | 1.470 | $(0.087)$ |
| 2 | 1.414 | $(0.089)$ |
| 3 | 1.413 | $(0.094)$ |
| 4 | 1.430 | $(0.099)$ |
| 5 | 1.566 | $(0.104)$ |
| 6 | 1.659 | $(0.105)$ |
| 7 | 1.608 | $(0.109)$ |
| 8 | 1.538 | $(0.113)$ |
| 9 | 1.495 | $(0.117)$ |
| 10 | 1.496 | $(0.121)$ |
| 11 | 1.518 | $(0.125)$ |
| 12 | 1.596 | $(0.129)$ |
| 13 | 1.693 | $(0.132)$ |
| 14 | 1.646 | $(0.135)$ |
|  | Underlying Hazard | Rate |
| $t$ | -0.048 | $(0.009)$ |
| $t^{2}$ | 0.0004 | $(0.0001)$ |
| Constant | -1.893 | $(0.107)$ |
|  |  |  |
|  |  |  |

Table A9: Marginal Effects on Probability of Discharge Other DRGs

| Day of stay $(t)$ | Probability of Discharge on Magic Day ${ }^{1}$ | Probability of Discharge on Day Preceding Magic Day ${ }^{2}$ | Hazard Ratio ${ }^{3}$ |
| :---: | :---: | :---: | :---: |
| DRG 189 |  |  |  |
| 19 | 11.02 | 1.73 | 6.39 |
|  | (0.358) | (0.074) | [204.9] |
| 20 | 11.40 | 1.81 | 6.29 |
|  | (0.353) | (0.080) | [203.3] |
| 21 | 11.77 | 1.90 | 6.20 |
|  | (0.352) | (0.086) | [201.3] |
| 22 | 12.11 | 1.98 | 6.12 |
|  | (0.354) | (0.093) | [199.0] |
| 23 | 12.23 | 2.06 | 6.05 |
|  | (0.358) | (0.101) | [196.5] |
| 24 | 12.72 | 2.13 | 5.98 |
|  | (0.364) | (0.109) | [193.8] |
| 25 | 12.99 | 2.19 | 5.92 |
|  | (0.372) | (0.117) | [191.0] |
| 26 | 13.23 | 2.25 | 5.87 |
|  | (0.382) | (0.125) | [188.0] |
| 27 | 13.43 | 2.30 | 5.83 |
|  | (0.393) | (0.134) | [184.9] |
| DRG 871 |  |  |  |
| 19 | 11.80 | 1.72 | 6.87 |
|  | (0.716) | (0.088) | [115.7] |
| 20 | 13.02 | 1.99 | 6.55 |
|  | (0.619) | (0.119) | [103.7] |
| 21 | 14.22 | 2.27 | 6.27 |
|  | (0.629) | (0.183) | [91.43] |
| DRG 177 |  |  |  |
| 19 | 9.56 | 2.54 | 3.77 |
|  | (0.499) | (0.120) | [86.05] |
| 20 | 10.22 | 2.77 | 3.69 |
|  | (0.567) | (0.139) | [91.55] |

Note: Standard errors in parentheses. P-values in brackets. This sample contains only episodes of hospitalization that terminated in discharge to home care or nursing facilities.
${ }^{1} \Phi\left(\gamma_{0}+\gamma_{1} t+\gamma_{2} t^{2}+\mu_{0}\right) * 100$
${ }^{2} \Phi\left(\gamma_{0}+\gamma_{1} t+\gamma_{2} t^{2}+\mu_{-1}\right) * 100$
${ }^{3}$ Hazard ratio: $\frac{\Phi\left(\gamma_{0}+\gamma_{1} t+\gamma_{2} t^{2}+\mu_{0}\right)}{\Phi\left(\gamma_{0}+\gamma_{1} t+\gamma_{2} t^{2}+\mu_{-1}\right)}$. Square brackets contain the p-value from a Wald test for $H_{0}: H R=\frac{\Phi\left(\gamma_{0}+\gamma_{1} t+\gamma_{2} t^{2}+\mu_{0}\right)}{\Phi\left(\gamma_{0}+\gamma_{1} t+\gamma_{2} t^{2}+\mu_{-1}\right)}=1$.

Table A10: Selected Probit Coefficients by Subgroup, DRG 207 at $d a y=29$

| Model \#/Partition | SSO Threshold Day | Preceding Day |
| :--- | :---: | :---: |
| Model \#1: |  |  |
| For-profit | 2.96 | 1.95 |
|  | $(0.332)$ | $(0.333)$ |
| Not for profit | 2.85 | 2.12 |
|  | $(0.340)$ | $(0.332)$ |
| Model \#2: |  |  |
| Kindred and Select | 3.09 | 2.05 |
|  | $(0.322)$ | $(0.324)$ |
| Other | 2.99 | 2.17 |
|  | $(0.328)$ | $(0.322)$ |
| Model \#3: |  |  |
| After Acquisition | 3.27 | 2.02 |
|  | $(0.247)$ | $(0.245)$ |
| Before Acquisition | 3.21 | 2.32 |
|  | $(0.254)$ | $(0.249)$ |
| Never Acquired | 3.13 | 2.22 |
|  | $(0.246)$ | $(0.242)$ |
| Model \#4: |  |  |
| HwH | 2.36 | 3.41 |
|  | $(0.284)$ | $(0.284)$ |
| Not HwH | 2.31 | 3.19 |
|  | $(0.282)$ | $(0.287)$ |

[^2]
## E Strategic Discharge and Capacity Constraints

This appendix considers whether capacity constraints affect LTCHs' decisions to strategically discharge patients. Measuring the capacity utilization of LTCHs is difficult because we only have data for their Medicare patients, and even then we only observe the quarter of the year in which their stays began. Nevertheless, we can create a rough proxy for capacity utilization by constructing a variable for each LTCH-quarter that gives the number of Medicare patient-days per LTCH bed per quarter across all the DRGs in our data. We use this variable to gauge how the probability of discharge varies across our measure of capacity utilization, with the idea being that hospitals that routinely have more patients per bed would also be more likely to be capacity constrained ${ }^{45}$ Table A11 shows the probability of being discharged on the magic day and the day before it for DRG 207, the main DRG we analyze in the paper. Broken down by decile based on the number of Medicare patient-days per LTCH bed per quarter (so the 10th decile is the set of hospitals that have the most Medicare patient-days per LTCH bed per quarter, i.e., the hospitals that are the most capacity constrained according to this rough measure of capacity utilization), the ratio of the discharge probabilities for the magic day over the day before it clearly shows that the probability of engaging in strategic discharge initially increases in the LTCH's capacity utilization, but then flattens out.

Table A11: Strategic Discharge by Capacity Utilization

| Medicare Patient Days | Probability of Discharge |  |  | P-value of difference |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Per Bed Decile | SSO Threshold Day | Preceding Day | Ratio | with lower decile |  |
| 1 | 0.048 | 0.016 | 2.99 | - |  |
| 2 | 0.066 | 0.013 | 5.23 | 0.000 |  |
| 3 | 0.078 | 0.011 | 7.29 | 0.000 |  |
| 4 | 0.081 | 0.01 | 8.17 | 0.250 |  |
| 5 | 0.09 | 0.01 | 8.8 | 0.001 |  |
| 6 | 0.088 | 0.009 | 9.79 | 0.417 |  |
| 7 | 0.095 | 0.009 | 10.79 | 0.022 |  |
| 8 | 0.091 | 0.01 | 9.39 | 0.219 |  |
| 9 | 0.091 | 0.011 | 8.65 | 0.980 |  |
| 10 | 0.09 | 0.01 | 8.76 | 0.866 |  |

Note: The proxy for capacity constraint is described in the text. Discharge results are for DRG 207.

Motivated by the summary statistics in Table A11, we further consider this issue using the same type of probit analysis as in Section 4.2. Table A12 shows that LTCHs in the first tercile of capacity utilization (i.e., the least capacity constrained) engage in less strategic discharging
than hospitals in the second or third tercile (between the second and third there is no statistically significant difference).
The next two tables show the heterogeneous effects of this relationship across different hospital
types. Table A13 repeats the analysis from Table A12 but interacts the extent of capacity utilization with indicators for whether the LTCH is a for-profit facility or not. The table shows that the difference in strategic discharging across for-profit and non-profit LTCHs is greatest

[^3]Table A12: Probit Marginal Effects by Capacity Utilization, DRG 207

|  | Predicted Prob. of Discharge |  | Hazard | Ratio of |
| :--- | :---: | :---: | :---: | :---: |
| Ratio | Hazard Ratios ${ }^{1}$ |  |  |  |
| Sercile of Capacity Utilization: |  |  |  |  |
| First | 7.36 | 1.31 | 5.63 |  |
|  | $(0.381)$ | $(0.121)$ | $[0.000]$ |  |
| Second | 9.38 | 1.00 | 9.31 | 1.65 |
|  | $(0.435)$ | $(0.079)$ | $[0.000]$ | $[0.000]$ |
| Third | 9.39 | 0.949 | 9.90 | 1.76 |
|  | $(0.481)$ | $(0.070)$ | $[0.000]$ | $[0.000]$ |

Note: Standard errors in parentheses. P-values in brackets. This sample contains only episodes of hospitalization that terminated in discharge to home care or nursing facilities.
${ }^{1}$ Ratio of hazard ratios are relative to the first tercile.
within the bottom tercile of capacity utilization. Furthermore, only non-profit LTCHs become more likely to strategically discharge patients as they become more capacity constrained. Table A14 repeats the analysis from Table A12 but interacts the extent of capacity utilization with indicators for whether the LTCH is owned by a chain or not. The table shows that chain-owned LTCHs engage in more strategic discharging than non-chain-owned LTCHs, but this difference is mainly at lower levels of capacity utilization. Finally, while chain-owned LTCHs do not increase their use of strategic discharge until they become very capacity constrained (third tercile), non-chain-owned LTCHs do so at lower levels of capacity utilization.

Table A13: Probit Marginal Effects by Capacity Utilization by For-Profit Status, DRG 207

|  | Predicted Prob. of Discharge |  | Hazard <br> Ratio | Within Ratio of Hazard Ratios ${ }^{1}$ | Across Ratio of Hazard Ratios ${ }^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | SSO Threshold Day | Preceding Day |  |  |  |
| Tercile of Capacity Utilization: |  |  |  |  |  |
| First: |  |  |  |  |  |
| For-Profit | 8.48 | 1.03 | 8.14 |  |  |
|  | (0.482) | (0.110) | [0.000] |  |  |
| Non-Profit | 5.54 | 1.79 | 3.09 |  | 2.63$[0.005]$ |
|  | (0.520) | (0.260) | [0.002] |  |  |
| Second: |  |  |  |  |  |
| For-Profit | 9.60 | 0.91 | 10.55 | 1.30 |  |
|  | (0.480) | (0.079) | [0.000] | [0.126] |  |
| Non-Profit | 8.76 | 1.29 | 6.79 | 2.19 | 1.55 |
|  | (0.730) | (0.200) | [0.000] | [0.026] | [0.104] |
| Third: ${ }^{\text {a }}$ (0.104 |  |  |  |  |  |
| For-Profit | 9.31 | 0.86 | 10.83 | 1.33 |  |
|  | (0.526) | (0.066) | [0.000] | [0.121] |  |
| Non-Profit | 9.79 | 1.39 | 7.04 | 2.28 | 1.54 |
|  | (1.130) | (0.215) | [0.000] | [0.048] | [0.146] |

Note: Standard errors in parentheses. P-values in brackets. This sample contains only episodes of hospitalization that terminated in discharge to home care or nursing facilities.
${ }^{1}$ Ratio of hazard ratios are relative to the first-tercile-capacity-row for the same type of hospital. That is, it compares for-profits (or non-profits) at different terciles of capacity utilization to those in the first tercile.
${ }^{2}$ Ratio of hazard ratios are relative to the same capacity-tercile-row for the other type of hospital. That is, it compares for-profits and non-profits within the same tercile of capacity utilization.

Table A14: Probit Marginal Effects by Capacity Utilization by Chain-Owned Status, DRG 207

|  | Predicted Prob. of Discharge |  | Hazard | Within Ratio of <br> Ratio | Across Ratio of <br> Hazard Ratios $^{1}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Hazard Ratios $^{2}$ |  |  |  |  |  |

Note: Standard errors in parentheses. P-values in brackets. This sample contains only episodes of hospitalization that terminated in discharge to home care or nursing facilities.
${ }^{1}$ Ratio of hazard ratios are relative to the first-tercile-capacity-row for the same type of hospital. That is, it compares Chain-owned (or Non-chain-owned) at different terciles of capacity utilization to those in the first tercile.
${ }^{2}$ Ratio of hazard ratios are relative to the same capacity-tercile-row for the other type of hospital. That is, it compares Chain-owned and Non-chain-owned within the same tercile of capacity utilization.

## F Additional Figures for Counterfactual Analysis

Figure A3 displays the observed (solid line) discharge probabilities over time and the predicted (dashed line) discharge probabilities corresponding to the estimates in column (1) of Table 5, where the horizontal axis in these figures is the number of days relative to the magic day (vertical line).$^{46}$ Panel (a) compares the predicted and observed discharge distributions for the entire sample of pooled DRGs while panel (b) focuses on just DRG 207.

Figure A3: Observed and predicted discharge probabilities


Figure A4 compares the reimbursement schemes we consider in the counterfactual analysis to the current PPS.

[^4]

Figure A4: Counterfactual Reimbursement Policies for DRG 207.


[^0]:    ${ }^{43}$ To discourage LTCHs from avoiding extremely high-cost patients, Medicare will share costs beyond what are reimbursed through the standard long-term care payment. In 2015, for example, if the costs incurred by an LTCH were more than the full long-term care payment plus $\$ 14,972$, then Medicare will pay 80 percent of the difference. According to our data, this happens in about 10 percent of long-term stays for DRG 207.

[^1]:    ${ }^{44}$ See http://www.cms.gov/Outreach-and-Education/Medicare-Learning-Network-MLN/
    MLNMattersArticles/downloads/MM10273.pdf for a detailed list of recent policy changes regarding LTCH payments.

[^2]:    Note: Standard errors in parentheses.

[^3]:    ${ }^{45}$ It should be noted that this proxy for capacity utilization may suffer from non-classical measurement error, as it may be correlated with other factors such as the Medicare share of total hospital days at each LTCH.

[^4]:    ${ }^{46}$ The predicted discharge probabilities are computed by simulating the model 100,000 times.

